

**ON ACCURACY OF  
APPROXIMATIONS FOR  
MULTIVARIATE SCALE  
MIXTURES IN STATISTICAL  
APPLICATIONS**

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## Introduction 1

# Some Problems on Asymptotic Expansions

Edgeworth Expansion

$$\begin{aligned} P(T \leq x) &= G_f(x) \\ &+ \frac{1}{n} \sum_{j=0}^k b_j G_{f+2j}(x) + R_2(x) \\ &= G_f(x) - \frac{1}{n} a(x) g_f(x) + R_2(x) \end{aligned}$$

Cornish-Fisher Expansion

$$\begin{aligned} P(T \leq t(u)) &= G_f(u) = 1 - \alpha \\ t(u) &= u + \frac{1}{n} a(u) + \tilde{R}_2(u) \end{aligned}$$

## Introduction 2

(i) **How to find asymptotic expansions**

(ii) **How to establish validity**

$$R_2(x) = O(n^{-2})$$

(iii) **How to find error bounds**

(1) Order estimate;  $R_2(x) = O(n^{-2})$

$\exists n_0, c$  s.t.  $|R_2(x)| \leq cn^{-2}$  for  $n \geq n_0$   
( $n_0, c$ ; unknown)

(2) Berry-Esseen type bounds  
with known constant  $c$

$|R_2(x)| \leq ch(\theta)n^{-2}$  for all  $n$   
( $h(\cdot)$  – known function)

## Introduction 3

# Berry-Esseen Bound

## Sum of i.i.d. variables

Let:

$$X_1, X_2, \dots, X_n \sim i.i.d.X$$

$$E(X) = 0 \quad E(X^2) = 1 \quad E(|X|^3) = \rho < \infty$$

$$\text{Let } Y_n = n^{-1/2} \sum_{j=1}^n X_j$$

Then for any  $x$  and  $n = 1, 2, \dots$

$$|P(Y_n \leq x) - \Phi(x)| \leq \frac{c\rho}{\sqrt{n}}$$

$c$  – absolute constant

$$c \leq 0.7655$$

## Introduction 4

### *t*-Statistic

Let:

$$X_1, X_2, \dots, X_n \sim i.i.d.X$$

$$E(X) = 0 \quad E(X^2) = \sigma^2 \quad E(|X|^3) = \rho < \infty$$

Let:  $T_n = \frac{\sqrt{n}\bar{X}}{\hat{\sigma}}, \quad n \geq 2$

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{j=1}^n (X_j - \bar{X})^2$$

Then for any  $x$  and  $n = 2, 3, \dots$

there exists an absolute constant  $c$  such that

$$|\mathbb{P}(T_n \leq x) - \Phi(x)| \leq \frac{c\rho}{\sqrt{n}}$$

Bentkus and Götze (1996)

## Statistic 1

# One-Way MANOVA

### Model

$$\mathbf{y}_j^{(i)} = \boldsymbol{\mu}^{(i)} + \boldsymbol{\varepsilon}_j^{(i)} \quad (j = 1, \dots, N_i) \\ i = 1, \dots, q + 1)$$

$$\boldsymbol{\varepsilon}_j^{(i)} : \text{Independent} \\ E(\boldsymbol{\varepsilon}_j^{(i)}) = 0, \text{Var}(\boldsymbol{\varepsilon}_j^{(i)}) = \boldsymbol{\Sigma}$$

### Hypothesis Testing

$$H_0 : \boldsymbol{\mu}^{(1)} = \dots = \boldsymbol{\mu}^{(q+1)}$$

VS

$$H_1 : \boldsymbol{\mu}^{(i)} \neq \boldsymbol{\mu}^{(j)} \text{ for some } i, j$$

### SS & SP Matrices

$$S_h = \sum_{i=1}^{q+1} N_i (\bar{\mathbf{y}}^{(i)} - \bar{\mathbf{y}})(\bar{\mathbf{y}}^{(i)} - \bar{\mathbf{y}})'$$

$$S_e = \sum_{i=1}^{q+1} \sum_{j=1}^{N_i} (\mathbf{y}_j^{(i)} - \bar{\mathbf{y}}^{(i)})(\mathbf{y}_j^{(i)} - \bar{\mathbf{y}}^{(i)})'$$

## Statistic 2

# Multivariate Linear Model

## Model

$$Y = X\Theta + (\varepsilon_1, \dots, \varepsilon_n)'$$

$$\varepsilon_1, \dots, \varepsilon_n \sim i.i.d.$$

$$E[\varepsilon_j] = 0, \quad \text{Cov}(\varepsilon_j) = \Sigma$$

## Hypothesis Testing

$$H_0 : C\Theta = 0$$

## SS & SP Matrices

$$S_h = \hat{\Theta}'C'\{C(X'X)^{-1}C'\}^{-1}C\hat{\Theta}$$

$$S_e = Y'(I_n - X(X'X)^{-1}X')Y$$

$$\hat{\Theta} = (X'X)^{-1}X'Y$$

## Statistic 3

### Test Statistics

$$(i) \quad T_{LR} = -n \log\{|S_e|/|S_e + S_h|\}$$

$$(ii) \quad T_{LH} = T_0^2 = n \operatorname{tr} S_h S_e^{-1}$$

$$S_e \sim W_p(n, \Sigma) \quad S_h \sim W_p(q, \Sigma)$$

$S_e \perp S_h$  independent;  $\Sigma = I$

The case  $p = 1$

$$(i) \quad T_{LR} = n \log \left( 1 + \frac{1}{n} T_0^2 \right)$$

$$(ii) \quad T_{LH} = T_0^2 = \left( \frac{1}{n} \chi_n^2 \right)^{-1} \chi_q^2$$

## Result 1

### Asymptotic Approximations for $P(T_{LR} \leq x)$ under Null Case

Asymptotic Expansion; Box (1949)

$$\begin{aligned}P(T_{LR} \leq x) &= G_r(x) + \frac{1}{n} \sum_{j=0}^1 b_j G_{r+2j}(x) \\ &\quad + O(n^{-2}) \\ &= G_r(x) - \frac{x}{2n} (q - p - 1) g_r(x) \\ &\quad + O(n^{-2})\end{aligned}$$

$$r = pq$$

$$b_0 = r(q - p - 1)/4$$

$$b_1 = -r(q - p - 1)/4$$

## Result 2

### Asymptotic Approximations for $P(T_0^2 \leq x)$ under Null Case

Asymptotic Expansion; Siotani (1956)

$$\begin{aligned} P(T_0^2 \leq x) &= G_r(x) + \frac{1}{n} \sum_{j=0}^2 b_j G_{r+2j}(x) \\ &\quad + O(n^{-2}) \\ &= G_r(x) \\ &\quad + \frac{x}{2n} \{q - p - 1 - (q + p + 1)x\} g_r(x) \\ &\quad + O(n^{-2}) \end{aligned}$$

$$\begin{aligned} b_0 &= r(q - p - 1)/4, & b_1 &= -r q/2, \\ b_2 &= r(q + p + 1)/4, & r &= pq \end{aligned}$$

### Result 3

## Error Bound for Asymptotic Expansions of $P(T_0^2 \leq x)$

**UFS (2003)**

$$|P(T_0^2 \leq x) - G_r(x)| \leq \frac{1}{2n} p(p+1) \nu_{-1,2,p}$$

$$\nu_{-1,2,p} = \eta_{-1,2} + \frac{1}{2}(p-1)\xi_{-1,1}\eta_{-1,1}$$

$$g_r(x) = \{2^{r/2}\Gamma(r/2)\}^{-1} \exp(-x/2)x^{r/2-1}$$

$$b_{-1,1}(x) = -\frac{1}{2}(x-r)$$

$$b_{-1,2}(x) = -\frac{1}{4}\{x^2 - 2rx + r(r-2)\}$$

$$\xi_{-1,1} = \int_{-\infty}^{\infty} |b_{-1,1}(x)g(x)|dx$$

$$\xi_{-1,2} = \frac{1}{2} \int_{-\infty}^{\infty} |b_{-1,2}(x)g(x)|dx$$

$$\eta_{-1,1} = \xi_{-1,1} + 2$$

$$\eta_{-1,2} = \{\sqrt{\xi_{-1,2}} + \sqrt{2 + \xi_{-1,1}}\}^2$$

## Result 4

### Values of $\nu_{-1,2,p}(\lambda)$ in gamma case

		$p$			
$\delta$	$\lambda$	1	2	3	4
-1	0.5	4.574	5.175	5.776	6.377
	1.0	5.711	6.718	7.724	8.731
	1.5	6.691	8.044	9.397	10.750
	2.0	7.608	9.277	10.946	12.615
	2.5	8.469	10.434	12.399	14.364
	3.0	9.286	11.534	13.782	16.030
	3.5	10.069	12.589	15.109	17.630
	4.0	10.825	13.609	16.394	19.178

## Result 5

### Error Bound for Asymptotic Expansions of $P(T_0^2 \leq x)$

$$\begin{aligned} & |P(T_0^2 \leq x) - G_r(x) - \frac{r}{4n} \{(q - p - 1)G_r(x) \\ & \quad - 2qG_{r+2}(x) + (q + p + 1)G_{r+4}(x)\}| \\ & \leq \frac{r}{48n^2} (|h_1| + |h_2| + 48q) \\ & \quad + \frac{1}{2n^2} p(2p^2 + 5p + 5) \min\{\eta_{-1,4,p}, \nu_{-1,4,p}\} \end{aligned}$$

$$\eta_{\delta,4,p} = \left\{ V_{\delta,k,p}^{1/4} + \left( 2 + p \sum_{j=1}^3 V_{\delta,j,p} \right)^{1/4} \right\}^4$$

## Result 6

### Coefficients

$$V_{\delta,1,p} = \xi_{\delta,1}$$

$$V_{\delta,2,p} = \xi_{\delta,2} + \frac{1}{2}(p-1)\xi_{\delta,1}^2$$

$$V_{\delta,3,p} = \xi_{\delta,3} + (p-1)\xi_{\delta,1}\xi_{\delta,2} + \frac{1}{6}(p-1)(p-2)\xi_{\delta,1}^3$$

$$\begin{aligned} V_{\delta,4,p} = & \xi_{\delta,4} + \frac{1}{2}\xi_{\delta,2}^2 + (p-1)\xi_{\delta,1}\xi_{\delta,3} \\ & + \frac{1}{2}(p-1)(p-2)\xi_{\delta,1}^2\xi_{\delta,2} \\ & + \frac{1}{24}(p-1)(p-2)(p-3)\xi_{\delta,1}^4 \end{aligned}$$

## Result 7

### Error Bound for Asymptotic Expansions of $P(T_{LR} \leq x)$

FU (2003)

$$\begin{aligned} & |P(T_{LR} \leq x) - G_r(x)| \\ & \leq \frac{1}{n} \{B_{-1,2,p} + C_{-1,2,p} + D_1\} \end{aligned}$$

$$B_{-1,2,p} = \nu_{-1,2,p} \left\{ 1 + \frac{1}{12n} (p-1)(2p-7) \right\}$$

$$C_{-1,2,p} = \frac{r}{4} (p-1)$$

$$\begin{aligned} D_1 = \frac{r}{4} & \left[ (4 + |q-2|) \left\{ 1 - n^{-1}q \right\}^{-(q+2)/2} \right. \\ & \left. + (q+2) \left\{ 1 - n^{-1}(q+2) \right\}^{-(q+4)/2} \right] \end{aligned}$$

## Result 8

### Error Bound for Asymptotic Expansions of $P(T_{LR} \leq x)$

$$\begin{aligned} & |P(T_{LR} \leq x) - G_r(x) \\ & \quad - \frac{r}{4n}(q - p - 1)\{G_r(x) - G_{r+2}(x)\}| \\ & \leq \frac{1}{n^2} \{B_{-1,4,p} + C_{-1,4,p} \\ & \quad + D_2 + D_3\} \end{aligned}$$

## Outline 1

### An Expression for $T_0^2 (= T_{LH})$

$$\begin{aligned} T_0^2 &= n \operatorname{tr}(U'U) S_e^{-1} \\ &= n \operatorname{tr}(H'U'UH) (H'S_eH)^{-1} \\ &\quad U = (u_{ij}); \quad u_{ij} \sim i.i.d.N(0, 1) \\ &= X_1 + \dots + X_p \end{aligned}$$

$$\mathbf{X} = (X_1, \dots, X_p)'$$

$$(i) \quad X_i = S_i Z_i, \quad i = 1, \dots, p$$

$$(ii) \quad Z_1, \dots, Z_p \sim i.i.d. \chi_q^2$$

$$(iii) \quad S_i = Y_i^{-1} (i = 1, \dots, p)$$

$$Y_1 > \dots > Y_p > 0;$$

the characteristic roots of  $W$

$$nW \sim W_p(n, I_p)$$

## Outline 2

### An Expression for $T_{LR}$

$$\begin{aligned}T_{LR} &= n \log \Lambda \\ &= n \log \prod_{j=1}^p V_j \\ & \quad V_j \sim Be\left(\frac{1}{2}(n-j+1), \frac{1}{2}q\right) \\ &= \sum_{i=1}^p n \log(1 + X_i/n)\end{aligned}$$

$$\mathbf{X} = (X_1, \dots, X_p)'$$

$$(i) \quad X_i = S_i Z_i, \quad i = 1, \dots, p$$

$$(ii) \quad Z_1, \dots, Z_p \sim i.i.d. \chi_q^2$$

$$(iii) \quad S_i = Y_i^{-1} (i = 1, \dots, p)$$
$$nY_i \sim \chi_{m_i}^2, \quad m_i = n - (i - 1)$$

## Outline 3

### Multivariate Scale Mixture

$$\mathbf{X} = (X_1, \dots, X_p)' \quad X_i = S_i Z_i, i = 1, \dots, p$$

- (i)  $Z_1, \dots, Z_p \sim i.i.d.$
- (ii)  $S_i > 0 \perp Z_i, i = 1, \dots, p$

Approximation for the density  $f(\mathbf{x})$  of  $\mathbf{X}$ ;  
 $\hat{f}(\mathbf{x}) = g_{\delta, k, p}(\mathbf{x})$

Error Bound:

$$\begin{aligned} & \left| \mathbb{P}(\mathbf{X} \in A) - \int_A g_{\delta, k, p}(\mathbf{x}) d\mathbf{x} \right| \\ & \leq c_{\delta, k, p} \sum_{i=1}^p \mathbb{E}[|Y_i - 1|^k] \\ & Y_i = S_i^{-1} \quad k = 2, 4 \end{aligned}$$

## Outline 4

### Approximation Formulas

$$P(T \in A) = Q_{\delta,k,p}(A) = \int_A g_{\delta,k,p}(\mathbf{x}) d\mathbf{x}$$

$$Q_{-1,2,p}(A) = Q_{-1,2,p}^{(0)}(A) + Q_{-1,2,p}^{(1)}(A)$$

$$Q_{-1,4,p}(A) = Q_{-1,2,p}(A) + Q_{-1,4,p}^{(2)}(A) \\ + Q_{-1,4,p}^{(3)}(A)$$

(1) The Case  $T_0^2$ :

$$A = \{(x_1, \dots, x_p) \in \mathbf{R}^p : \\ x_1 + \dots + x_p \leq x\}$$

(2) The Case  $T_{LR}$ :

$$A = \tilde{A} \\ = \{(x_1, \dots, x_p) \in \mathbf{R}^p : \\ \sum_{i=1}^p n \log(1 + x_i/n) \leq x\}$$

## Outline 5

### Reduction When $A$ Is Symmetric

$$Q_{\delta,k,p}(A) = \int_A g_{\delta,k,p}(\mathbf{x})d\mathbf{x} = \sum_{j=0}^{k-1} Q^{(j)}(A)$$

$$Q^{(0)}(A) = \int_A g_p(\mathbf{x})d\mathbf{x}$$

For  $j \geq 1$ ;

$$\begin{aligned} Q^{(j)}(A) &= \sum_{(j)} \frac{1}{j_1! \dots j_p!} \\ &\times \int_A b_{\delta,j_1}(x_1) \dots b_{\delta,j_p}(x_p) g_p(\mathbf{x})d\mathbf{x} \\ &\times \mathbb{E} \left[ (Y_1 - 1)^{j_1} \dots (Y_p - 1)^{j_p} \right]. \end{aligned}$$

$$Q^{(1)}(A) = \int_A b_{\delta,1}(x_1) g_p(\mathbf{x})d\mathbf{x} \times \text{tr}V$$

$$\begin{aligned} Q^{(2)}(A) &= \frac{1}{2} \int_A b_{\delta,2}(x_1) g_p(\mathbf{x})d\mathbf{x} \times \text{tr}V^2 \\ &+ \int_A b_{\delta,1}(x_1) b_{\delta,1}(x_2) g_p(\mathbf{x})d\mathbf{x} \\ &\times \{(\text{tr}V)^2 - \text{tr}V^2\} \end{aligned}$$

$$V = S - I \quad S = \text{diag}(S_1, \dots, S_p)$$

# Univariate Scale Mixture 1

## $L_1$ -Norm Error Bound-Univariate

$$X = SZ \quad S > 0 \perp Z, \quad S = Y^{\delta\rho}$$

Let  $g$  be density function of  $Z$

Then density function of  $X$ :

$$f(x) = \mathbb{E}[Y^{-\delta\rho} g(xY^{-\delta\rho})]$$

Functions  $b_{\delta,j}$ :

$$\left. \frac{\partial^j}{\partial y^j} \left( y^{-\delta\rho} g(xy^{-\delta\rho}) \right) \right|_{y=1} = b_{\delta,j}(x)g(x)$$

Approximation for  $y^{-\delta\rho} g(xy^{-\delta\rho})$ ;

$$g_{\delta,k}(x, y) = g(x) + \sum_{j=1}^{k-1} \frac{1}{j!} b_{\delta,j}(x)g(x)(y-1)^j$$

Approximation for  $f(x)$ ;

$$g_{\delta,k}(x) = \mathbb{E}[g_{\delta,k}(x, Y)]$$

## Univariate Scale Mixture 2

### Main Result

$$\begin{aligned} & \left| \mathbb{P}(X \in A) - \int_A g_{\delta,k}(x) dx \right| \\ & \leq \frac{1}{2} \eta_{\delta,k} \mathbb{E}[|Y - 1|^k] \end{aligned}$$

$$\eta_{\delta,k} = \left\{ \xi_{\delta,k}^{1/k} + \left( 2 + \sum_{j=1}^{k-1} \xi_{\delta,j} \right)^{1/k} \right\}^k$$

$$\begin{aligned} \xi_{\delta,j} &= \frac{1}{j!} \left\| b_{\delta,j}(x) g(x) \right\|_1 \\ &= \frac{1}{j!} \int_{-\infty}^{\infty} |b_{\delta,j}(x) g(x)| dx \end{aligned}$$

# Multivariate Scale Mixture 1

## Multivariate Scale Mixture

$$\mathbf{X} = (X_1, \dots, X_p)' \quad X_i = S_i Z_i, i = 1, \dots, p$$

- (i)  $Z_1, \dots, Z_p \sim i.i.d.$   
(ii)  $S_i > 0 \perp Z_i, i = 1, \dots, p$

Transformation;

$$Y_i = S_i^{\delta/\rho} \quad \delta = \pm 1 \quad \rho = 1, 1/2$$

The Conditional Density of  $\mathbf{X}$   
given  $Y_i = y_i, i = 1, \dots, p$

$$y_1^{-\delta\rho} g(xy_1^{-\delta\rho}) \dots y_p^{-\delta\rho} g(xy_p^{-\delta\rho})$$

$$\begin{aligned} & g_{\delta,k,p}(\mathbf{x}) \\ &= \mathbb{E} \left[ g_p(\mathbf{x}) + \sum_{j=1}^{k-1} \frac{1}{j!} \left( (Y_1 - 1) \frac{\partial}{\partial y_1} + \dots + (Y_p - 1) \frac{\partial}{\partial y_p} \right)^j \right. \\ & \quad \left. \times y_1^{-\delta\rho} g(x_1 y_1^{-\delta\rho}) \dots y_p^{-\delta\rho} g(x_p y_p^{-\delta\rho}) \Big|_{y_1=\dots=y_p=1} \right] \\ &= g_p(\mathbf{x}) + \sum_{j=1}^{k-1} \sum_{(j)} \frac{1}{j_1! \dots j_p!} b_{\delta,j_1}(x_1) \dots b_{\delta,j_p}(x_p) g_p(\mathbf{x}) \\ & \quad \times \mathbb{E} [(Y_1 - 1)^{j_1} \dots (Y_p - 1)^{j_p}] \end{aligned}$$

## Multivariate Scale Mixture 2

### Main Result (I)

$$\begin{aligned} & \left| \mathbb{P}(\mathbf{X} \in A) - \int_A g_{\delta,k,p}(\mathbf{x}) d\mathbf{x} \right| \\ & \leq \frac{1}{2} \eta_{\delta,k,p} \sum_{i=1}^p \mathbb{E}[|Y_i - 1|^k]. \end{aligned}$$

$$\begin{aligned} \eta_{\delta,1,p} &= 2 + V_{\delta,1,p} \\ \text{For } k &\geq 2 \end{aligned}$$

$$\eta_{\delta,k,p} = \left\{ V_{\delta,k,p}^{1/k} + \left( 2 + p \sum_{j=1}^{k-1} V_{\delta,j,p} \right)^{1/k} \right\}^k,$$

$$V_{\delta,j,p} = \sum_{[j]} \frac{(p-1)!}{i_1! \dots i_m!} \xi_{\delta,j_1} \dots \xi_{\delta,j_p},$$

$$V_{\delta,1,p} = \xi_{\delta,1}$$

$$V_{\delta,2,p} = \xi_{\delta,2} + \frac{1}{2}(p-1)\xi_{\delta,1}^2$$

## Multivariate Scale Mixture 3

### Main Result (II)

$$\begin{aligned} & \left| \mathbb{P}(\mathbf{X} \in A) - \int_A g_{\delta,k,p}(\mathbf{x}) d\mathbf{x} \right| \\ & \leq \frac{1}{2} \nu_{\delta,k,p} \sum_{i=1}^p \mathbb{E}[|Y_i - 1|^k]. \end{aligned}$$

For  $k \geq 2$ ;

$$\nu_{\delta,k,p} = p^{-1} \left( \eta_{\delta,k} + (p-1) \sum_{q=0}^{k-1} \nu_{\delta,k-q,p-1} \xi_{\delta,q} \right)$$

For  $k \geq 1$ ;

$$\nu_{\delta,1,p} = \eta_{\delta,1}, \quad \nu_{\delta,k,0} = 0$$

$$\nu_{\delta,k,1} = \eta_{\delta,k}$$

$$\nu_{\delta,1,p} = \eta_{\delta,1},$$

$$\nu_{\delta,2,p} = \eta_{\delta,2} + \frac{1}{2}(p-1)\xi_{\delta,1}\eta_{\delta,1}$$

$$\eta_{\delta,1,p} = \nu_{\delta,1,p} \quad \eta_{\delta,2,p} \geq \nu_{\delta,2,p}$$