

# Fluctuations and Extreme Values in Multifractal Patterns<sup>1</sup>

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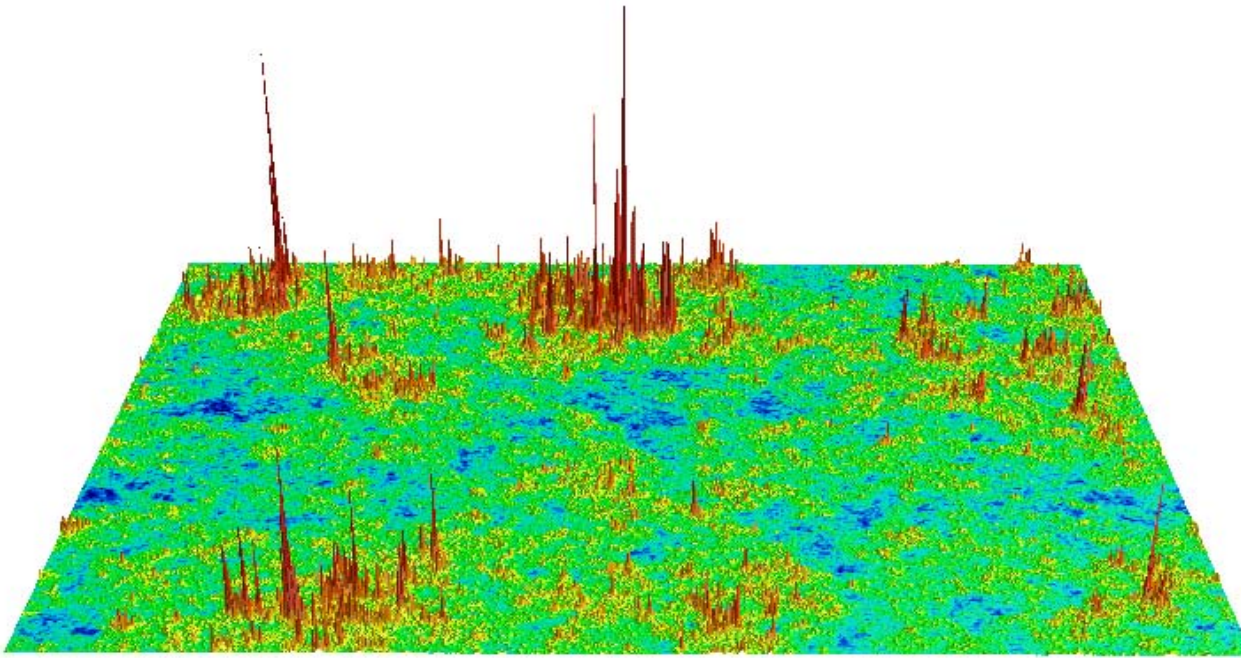
**D.I. Diakonov memorial symposium, St. Petersburg, 13/08/2013**

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<sup>1</sup>Based on: **YVF**, **P Le Doussal** and **A Rosso** J Stat Phys: **149** (2012), 898-920

## Disorder-generated multifractals:

Disorder-generated multifractal patterns display high variability over a wide range of space or time scales, associated with huge fluctuations in intensity which can be visually detected. Another common feature is presence of certain long-ranged **powerlaw-type correlations** in data values.

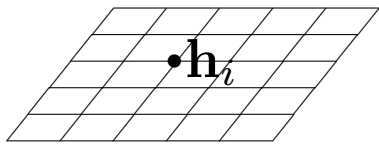


Intensity of a multifractal wavefunction at the point of Integer Quantum Hall Effect.

Courtesy of F. Evers, A. Mirlin and A. Mildenberger.

## Multifractal Ansatz:

Consider a certain (e.g. hypercubic) lattice of linear extent  $L$  and lattice spacing  $a$  in  $d$ -dimensional space, with  $M \sim (L/a)^d \gg 1$  being the total number of sites in the lattice. The multifractal patterns are then usually associated with a set of non-negative "heights"  $h_i \geq 0$  attributed to every lattice site  $i = 1, 2, \dots, M$  such that the heights scale in the limit  $M \rightarrow \infty$  differently at different sites:

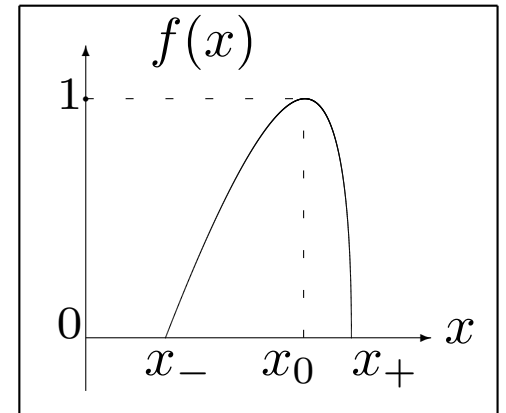


$$h_i \geq 0, \quad h_i \sim M^{x_i}$$

with exponents  $x_i$  forming a dense set such that

$$\rho_M(x) = \sum_{i=1}^M \delta \left( \frac{\ln h_i}{\ln M} - x \right) \approx c_M(x) \sqrt{\ln M} M^{f(x)}$$

We will refer below to the above form of the density as the **multifractal Ansatz**.



The major effort in the last decades was directed towards determining the shape and properties of the **singularity spectrum** function  $f(x)$ . In contrast, our main object of interest will be understanding the sample-to-sample fluctuations of the prefactor  $c_M(x)$  in disorder-generated multifractal patterns like those in the field of **Anderson localization**. Such fluctuations are reflected in statistics of the number of lattice points  $i$  satisfying  $h_i > M^x$  which is given by the **counting function**

$$\mathcal{N}_M(x) = \int_x^\infty \rho_M(y) dy.$$

## Thermodynamic formalism for multifractals:

When dealing with multifractal patterns it is frequently more convenient to characterize them by the set of exponents  $\zeta_q$  describing the large- $M$  scaling behaviour of the so-called **partition functions**

$$Z_q = \sum_{i=1}^M h_i^q \sim M^{\zeta_q}, \quad \ln M \gg 1$$

The counting function  $\mathcal{N}_M(x)$  is related to the partition function  $Z_q$  via the common density  $\rho_M(y)$  as

$$\mathcal{N}_M(x) = \int_x^\infty \rho_M(y) dy, \quad Z_q = \int_{-\infty}^\infty M^{qy} \rho_M(y) dy$$

Substituting for  $\rho_M(y)$  the **multifractal Ansatz** form  $\rho_M(y) \approx c_M(y) \sqrt{\ln M} M^{f(y)}$  we get for  $\ln M \gg 1$

$$\mathcal{N}_M(x) \approx \frac{c_M(x)}{|f'(x)| \sqrt{\ln M}} M^{f(x)}, \quad Z_q \approx c_M(y_*) \sqrt{\frac{2\pi}{|f''(y_*)|}} M^{\zeta_q}$$

where  $f'(y_*) = -q$  and  $\zeta_q = f(y_*) + q y_*$  are related by the **Legendre transform**.

**Conclusion:** The fluctuation properties of the **counting function**  $\mathcal{N}_M(x)$  and the **partition function**  $Z_q$  can be related to each other via the statistics of the common prefactor  $c_M(x)$ .

## From disorder-generated multifractals to log-correlated fields:

As discussed e.g. in **Duplantier, Ludwig** 1991 disorder-generated multifractal patterns of intensities  $h(\mathbf{r})$  are typically **self-similar**

$$\mathbb{E} \{h^q(\mathbf{r}_1)h^s(\mathbf{r}_2)\} \propto \left(\frac{L}{a}\right)^{y_{q,s}} \left(\frac{|\mathbf{r}_1 - \mathbf{r}_2|}{a}\right)^{-z_{q,s}}, \quad q, s \geq 0, \quad a \ll |\mathbf{r}_1 - \mathbf{r}_2| \ll L$$

and **spatially homogeneous**

$$\mathbb{E} \{h^q(\mathbf{r}_1)\} = \frac{1}{M} Z_q, \quad \text{with } M = \left(\frac{L}{a}\right)^d \text{ and } Z_q = \sum_{\mathbf{r}} h^q(\mathbf{r}) \propto \left(\frac{L}{a}\right)^{d\zeta_q}$$

The consistency of the two conditions for  $|\mathbf{r}_1 - \mathbf{r}_2| \sim a$  and  $|\mathbf{r}_1 - \mathbf{r}_2| \sim L$  implies:

$$y_{q,s} = d(\zeta_{q+s} - 1), \quad z_{q,s} = d(\zeta_{q+s} - \zeta_q - \zeta_s + 1)$$

If we now introduce the field  $V(\mathbf{r}) = \ln h(\mathbf{r}) - \mathbb{E} \{\ln h(\mathbf{r})\}$  and exploit the identities  $\frac{d}{ds} h^s|_{s=0} = \ln h$  and  $\zeta_0 = 1$  we arrive at the relation:

$$\mathbb{E} \{V(\mathbf{r}_1)V(\mathbf{r}_2)\} = -g^2 \ln \frac{|\mathbf{r}_1 - \mathbf{r}_2|}{L}, \quad g^2 = d \frac{\partial^2}{\partial s \partial q} \zeta_{q+s}|_{s=q=0}$$

**Conclusion:** logarithm of a multifractal intensity is a **log**-correlated random field.

**Strategy:** To understand statistics of **high values** and **positions of extremes** of general logarithmically correlated random processes and fields we first consider the simplest case of such process which is the **Gaussian**  $1/f$  **noise**.

## Ideal Gaussian periodic 1/f noise:

We will use a (regularized) model for ideal Gaussian periodic **1/f** noise defined as

$$V(t) = \sum_{n=1}^{\infty} \frac{1}{\sqrt{n}} [v_n e^{int} + \bar{v}_n e^{-int}] , \quad t \in [0, 2\pi)$$

where  $v_n, \bar{v}_n$  are **complex standard Gaussian i.i.d.** with  $\mathbb{E}\{v_n \bar{v}_n\} = 1$ . It implies the formal covariance structure:

$$\mathbb{E}\{V(t_1)V(t_2)\} = -2 \ln |2 \sin \frac{t_1 - t_2}{2}|, \quad t_1 \neq t_2$$

**Regularization procedure (YVF & Bouchaud 2008):** subdivide the interval  $[0, 2\pi)$  by a finite number  $M$  of observation points  $t_k = \frac{2\pi}{M}k$  where  $k = 1, \dots, M$ , and replace the function  $V(t), t \in [0, 2\pi)$  with a sequence of  $M$  random mean-zero Gaussian variables  $V_k$  correlated according to the  $M \times M$  **covariance matrix**  $C_{km} = \mathbb{E}\{V_k V_m\}$  such that the off-diagonal entries are given by

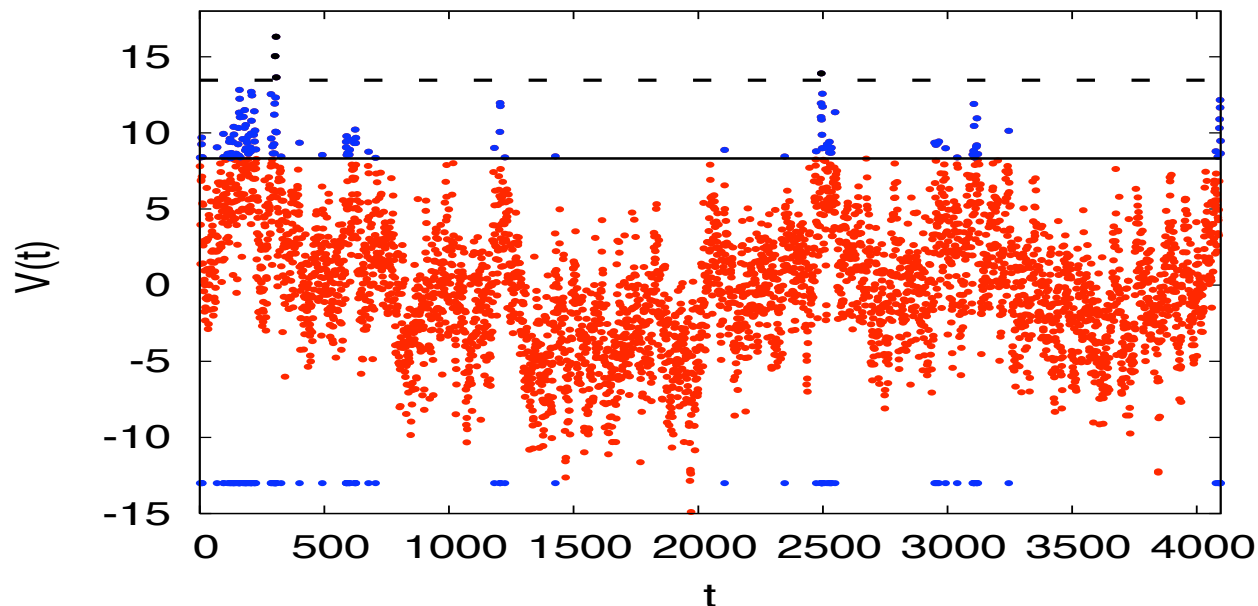
$$C_{k \neq m} = -2 \ln |2 \sin \frac{\pi}{M}(k - m)|, \quad C_{kk} = \mathbb{E}\{V_k^2\} > 2 \ln M, \quad \forall k = 1, \dots, M$$

The model is well defined, and we will actually take  $C_{kk} = 2 \ln M + \epsilon, \forall k$  with  $\epsilon \ll 1$ . We expect that the statistical properties of the sequence  $V_k$  generated in this way reflect for  $M \rightarrow \infty$  correctly the universal features of the  $1/f$  noise.

The **multifractal pattern** of heights is then generated by setting  $h_i = e^{V_i}$  for each  $i = 1, \dots, M$ .

## Circular-logarithmic model (YF & Bouchaud 2008):

An example of the  $1/f$  signal sequence generated for  $M = 4096$  according to the above prescription is given in the figure.



The upper line marks the typical value of the **extreme value threshold**  $V_m = 2 \ln M - \frac{3}{2} \ln \ln M$ .

The lower line is the level  $\frac{1}{\sqrt{2}}V_m$  and blue dots mark points supporting  $V_i > \frac{1}{\sqrt{2}}V_m$ .

**Questions we would like to answer:** How many points are typically above a given level of the noise? How strongly does this number fluctuate for  $M \rightarrow \infty$  from one realization to the other? How to understand the typical position  $V_m$  and statistics of the **extreme values** (maxima or minima), etc. And, after all, what parts of the answers are **universal** and what is the universality class?

## Characteristic polynomial of random CUE matrix and periodic 1/f noise:

Let  $U_N$  be a  $N \times N$  **unitary matrix**, chosen at random from the unitary group  $\mathcal{U}(N)$ . Introduce its **characteristic polynomial**  $p_N(\theta) = \det(1 - U_N e^{-i\theta})$  and further consider  $V_N(\theta) = -2 \log |p_N(\theta)|$ . Following **Hughes, Keating & O'Connell** 2001 one can employ the following representation

$$V_N^{(U)}(\theta) = -2 \log |p_N(\theta)| = \sum_{n=1}^{\infty} \frac{1}{\sqrt{n}} \left[ e^{-in\theta} v_n^{(N)} + \text{comp. conj.} \right]$$

where  $v_n^{(N)} = \frac{1}{\sqrt{n}} \text{Tr}(U_N^{-n})$ .

According to **Diaconis & Shahshahani** 1994 the coefficients  $v_n^{(N)}$  for any fixed  $n$  tend in the limit  $N \rightarrow \infty$  to i.i.d. complex gaussian variables with zero mean and variance  $\mathbb{E}\{|\zeta_n|^2\} = 1$ . We conclude that for finite  $N$  **Log-Mod** of the characteristic polynomial of CUE matrices is just a **certain regularization** of the stationary random **Gaussian Fourier series** of the form

$$V(t) = \sum_{n=1}^{\infty} \frac{1}{\sqrt{n}} \left[ v_n e^{int} + \bar{v}_n e^{-int} \right], \quad t \in [0, 2\pi)$$

where  $v_n, \bar{v}_n$  are **complex standard Gaussian i.i.d.** with  $\mathbb{E}\{v_n \bar{v}_n\} = 1$ .

**Random characteristic polynomials provide natural models for 1/f noise!**



**The distribution of the partition function**  $Z_q = \sum_{i=1}^M h_i^q$ ,  $h_i = e^{V_i}$ :

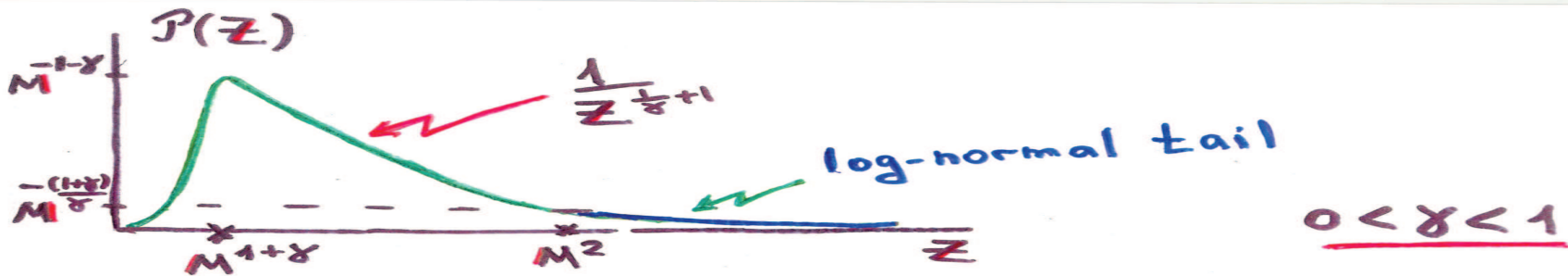
For  $M \gg 1$  the positive integer moments  $\mathbb{E}\{Z_q^n\}$  for  $q^2 < 1$  are given by

$$\mathbb{E}\{Z_q^n\} |_{M \gg 1} \approx \begin{cases} e^{n \ln M(1+q^2)} \mathcal{S}_n(q^2), & n < 1/q^2 \\ e^{\ln M(1+n^2q^2)} O(1), & n > 1/q^2 \end{cases}$$

where

$$\begin{aligned} \mathcal{S}_n(q^2) &= \frac{n!}{\pi^n} \int_0^\pi d\theta_1 \int_{\theta_1}^\pi d\theta_2 \int_{\theta_2}^\pi d\theta_3 \dots \int_{\theta_{n-1}}^\pi d\theta_n \prod_{p < q} [2 \sin(\theta_p - \theta_q)]^{-2q^2} \\ &= \frac{\Gamma(1-nq^2)}{\Gamma^n(1-q^2)} \quad \text{for } 1 < n < 1/q^2 \quad - \text{Dyson-Morris-Selberg integral.} \end{aligned}$$

One can use the above moments to restore the shape of the probability density  $\mathcal{P}_q(Z)$  for the partition function  $Z = Z_q$  in the whole domain  $\gamma = q^2 < 1$ :



$$\mathcal{P}_\gamma(Z) = \frac{1}{\gamma Z_e} \left(\frac{Z_e}{Z}\right)^{1+\frac{1}{\gamma}} e^{-\left(\frac{Z_e}{Z}\right)^{\frac{1}{\gamma}}}, \quad Z \sim Z_e = \frac{M^{1+\gamma}}{\Gamma(1-\gamma)} \ll M^2$$

$$\mathcal{P}_\gamma(Z) = \frac{M}{\sqrt{4\pi\gamma \ln M}} \frac{1}{Z} F\left(\frac{1}{2} \frac{\ln Z}{\ln M}\right) e^{-\frac{1}{4 \ln M \gamma} \ln^2 Z}, \quad Z \gg M^2$$

where  $F(x) \sim 1$  for  $x \sim 1$ .

## Statistics of the counting function $\mathcal{N}_M(x)$ and threshold of extreme values:

Applying the thermodynamic formalism in our particular case we conclude that the probability density for the (scaled) counting function  $n = \mathcal{N}_M(x)/\mathcal{N}_t(x)$  is given by:

$$\mathcal{P}_x(n) = \frac{4}{x^2} n^{-\left(1+\frac{4}{x^2}\right)} e^{-n^{-\frac{4}{x^2}}}, \quad n \ll n_c, \quad 0 < x < 2.$$

with  $n_c \rightarrow \infty$  for  $M \rightarrow \infty$  and the **characteristic scale**  $\mathcal{N}_t(x)$  given by

$$\mathcal{N}_t(x) = \frac{M^{f(x)}}{x\sqrt{\pi \ln M} \Gamma(1-x^2/4)} \text{ with the singularity spectrum } f(x) = 1 - x^2/4.$$

In particular, the position  $x_m$  of the **threshold of extreme values** is determined from the condition  $\mathcal{N}_t(x) \sim 1$ . This results in

$$x_m = 2 - c \frac{\ln \ln M}{\ln M} + O(1/\ln M) \text{ with } c = 3/2.$$

Note that  $\mathcal{N}_t(x) = \mathbb{E} \{ \mathcal{N}_M(x) \} \frac{1}{\Gamma(1-x^2/4)}$ . Had we instead decided to use the condition  $\mathbb{E} \{ \mathcal{N}_M(x) \} \sim 1$  that would give  $x_m = 2 - c \frac{\ln \ln M}{\ln M} + O(1/\ln M)$  with  $c = 1/2$ . The latter value is typical for **short-ranged** random sequences. The difference is due to the fact that for  $x \rightarrow 2$  the **typical** value  $\mathcal{N}_t(x)$  of the counting function is parametrically smaller than the **mean** value  $\mathbb{E} \{ \mathcal{N}_M(x) \}$ . We conjecture that this mechanism is common to all **logarithmically-correlated** processes and fields.

## Distribution of the absolute maximum: partition function approach:

Given the sequence  $\{V_i, i = 1, \dots, M\}$  we are interested in finding the **distribution** of  $V_{(m)} = \max(V_1, \dots, V_M)$  that is

$$P(v) = \text{Prob}(V_{(m)} < v) = \text{Prob}(V_i < v, \forall i) = \mathbb{E} \left\{ \prod_{i=1}^M \theta(v - V_i) \right\}$$

Next we use: 
$$\lim_{q \rightarrow \infty} \exp \left[ -e^{-q(v-V_i)} \right] = \begin{cases} 1 & v > V_i \\ 0 & v < V_i \end{cases} \equiv \theta(v - V_i)$$

which immediately shows that:

$$P(v) = \text{Prob}(V_{(m)} < v) = \lim_{q \rightarrow \infty} \mathbb{E} \left\{ \exp \left[ -e^{-qv} Z_q \right] \right\}, \quad \text{where } Z_q = \sum_{i=1}^M e^{qV_i}$$

From our previous knowledge of statistics of  $Z_q$  we can readily extract the function  $G_q(v) = \mathbb{E} \left\{ \exp \left[ -e^{-qv} Z_q \right] \right\}$  for  $q < 1$ . In the limit  $\ln M \gg 1$  that function turns out to be of the form:

$$G_q(v) = g_q(v - c_q \ln M) \quad \text{where } c_q = \left( q + \frac{1}{q} \right) \quad \text{and } g_q(v) = \int_0^\infty dt \exp \left\{ -t - e^{-qv} t^{-q^2} \right\}$$

One may further notice that not only  $c_q = c_{q-1}$  but the whole function satisfies a quite remarkable **duality relation**

$$g_q(v) = 1 + \sum_{n=1}^{\infty} \frac{(-1)^n}{n!} \left[ e^{-nqv} \Gamma(1 - nq^2) + e^{-n\frac{v}{q}} \Gamma\left(1 - \frac{n}{q^2}\right) \right] = g_{\frac{1}{q}}(x)$$

**THIS HOWEVER STILL DOES NOT ALLOW TO CONTINUE TO  $q > 1$ !**

## Freezing conjecture and the distribution of extremes:

Using certain analogy with the **Derrida-Spohn** model of polymers on disordered trees we conjecture the following **freezing scenario**: for the  $\log$ -circular model the same sort of **freezing transition** takes place at  $q = 1$ . Namely, the function

$$g_{q<1}(v) = \int_0^\infty dt \exp \left\{ -t - e^{-qv} t^{-q^2} \right\}$$

**freezes** to the **q-independent** profile  $g_{q=1}(v) = 2e^{-v/2} K_1(2e^{-v/2})$  in the whole "glassy" phase  $q > 1$ .

### **Consequences:**

(i) The latter profile then is precisely the distribution  $P(v)$  of the (shifted) absolute maximum:  $V_m = 2 \ln M - \frac{3}{2} \ln \ln M + v$ . This distribution is manifestly **non-Gumbel**, and shows the tail behaviour:  $P(v \rightarrow -\infty) \approx 1 - |v|e^v$

(ii) The probability density of the partition function  $Z_q$  in the whole regime  $q > 1$  must display a power-law forward tail of the form:

$$\mathcal{P}_{q>1}(Z) \propto Z^{-(1+\frac{1}{q})} \ln Z$$

This shape, including the meaningful **log-factor**, is believed to be **universal** for the whole class of logarithmically correlated processes.

## Numerics for the maxima of CUE characteristic polynomials:

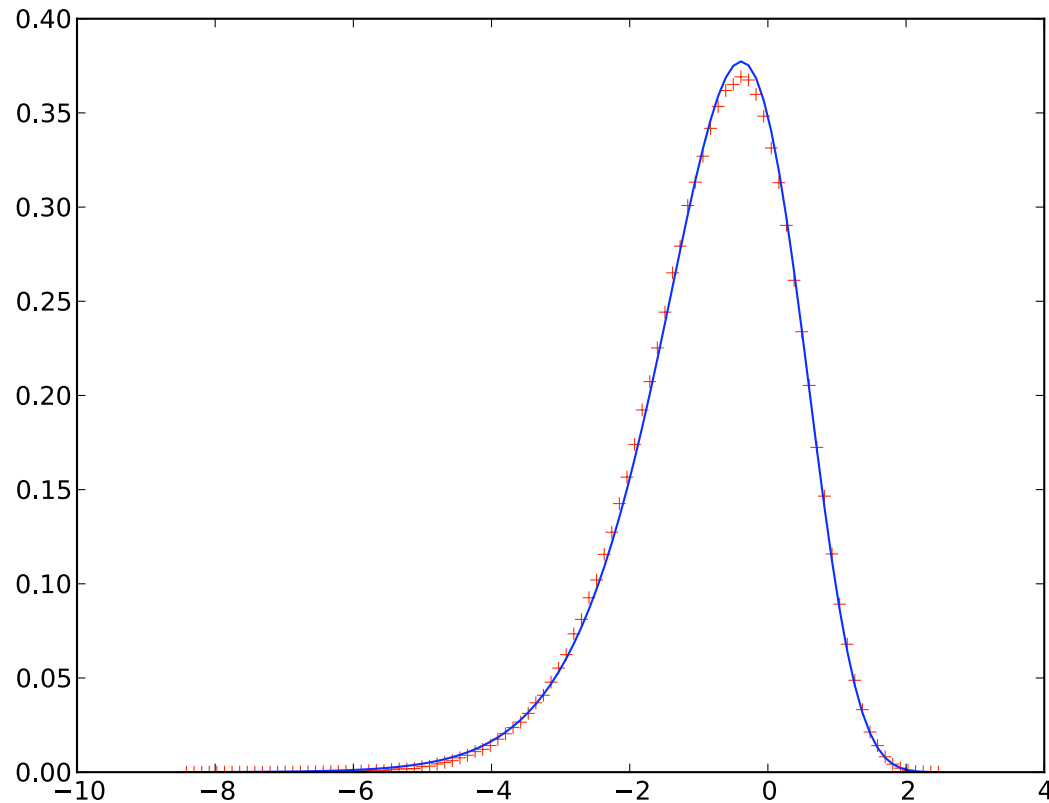


Figure 1: Density of maxima for CUE polynomials (  $N = 50, 10^6$  samples ) compared to periodic  $1/f$  noise prediction  $p(v) = 2e^v K_0(2e^{v/2})$ .

## Threshold of extreme values for self-similar multifractal fields:

**Conjecture:** the value  $c = \frac{3}{2}$  is a universal feature of systems with **logarithmic** correlations.

Apart from  $1/f$  noise and its incarnations (characteristic polynomials of random matrices, Riemann zeta-function along the critical line, and random Young diagrams sampled with the Plancherel measure) the new universality class is believed to include the  $2D$  Gaussian free field, branching random walks & polymers on disordered trees, some models in turbulence and financial mathematics and, with due modifications the **disorder-generated multifractals**.

Namely, consider a multifractal random **probability measure**  $p_i \sim M^{-\alpha_i}$ ,  $i = 1, \dots, M$  such that  $\sum_{i=1}^M p_i = 1$  characterized by a general non-parabolic **singularity spectrum**  $f(\alpha)$  with the left endpoint at  $\alpha = \alpha_- > 0$ . Then very similar consideration based on insights from **Mirlin & Evers** 2000 suggests that the **extreme value threshold** should be given by  $p_m = M^{-\alpha_m}$ , where  $\alpha_m$  is given by

$$\alpha_m \approx \alpha_- + \frac{3}{2} \frac{1}{f'(\alpha_-)} \frac{\ln \ln M}{\ln M} \quad \Rightarrow \quad -\ln p_m \approx \alpha_- \ln M + \frac{3}{2} \frac{1}{f'(\alpha_-)} \ln \ln M$$

For branching random walks this was recently rigorously proved: **L. Addario-Berry & B. Reed** 2009; **E. Aidekon** 2012

## Threshold of extreme values for self-similar multifractal fields:

**Work in progress:** testing such a prediction for multifractal eigenvectors of a  $N \times N$  random matrix ensemble introduced by E. Bogomolny & O. Giraud, *Phys. Rev. Lett.* **106** 044101 (2011) based on **Rujsenaars-Schneider** model of  $N$  interacting particles. Preliminary numerics is supportive of the theory.

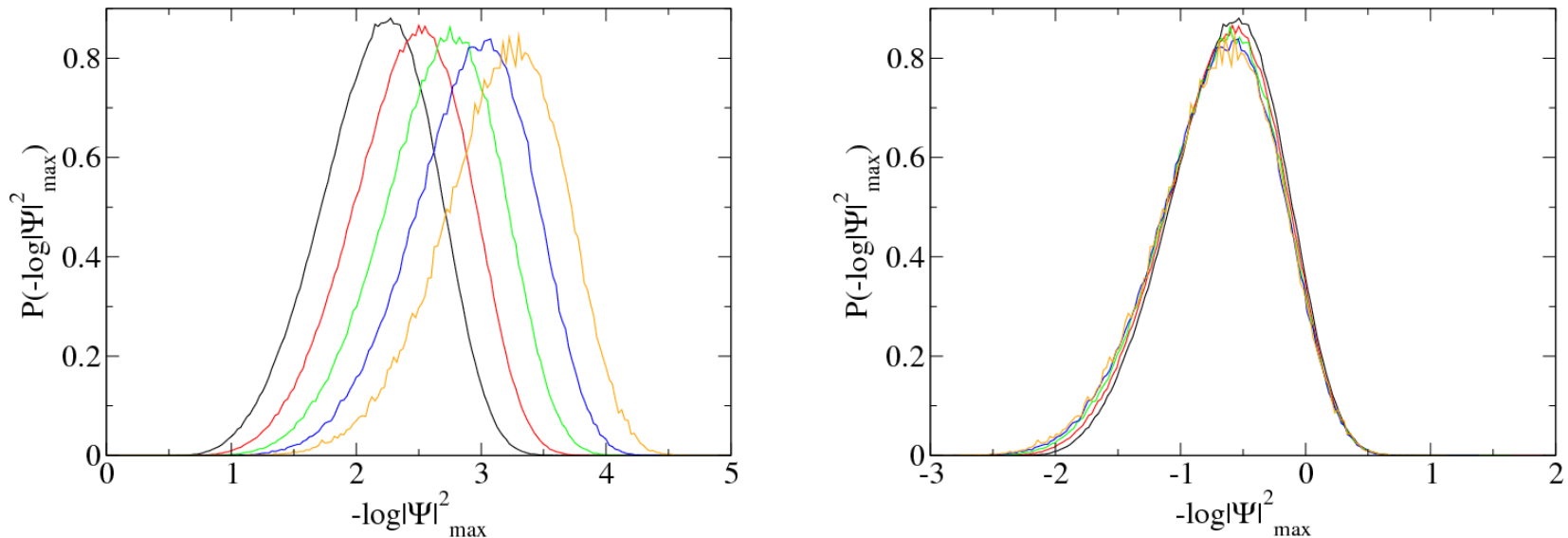


Figure 2: Statistics of maxima for eigenvectors of RS model for sample sizes  $M = 2^n$  with  $n = 8, \dots, 12$ . **left:** raw data **right:** each curve is shifted by  $\alpha_- \ln M + \frac{3}{2} \frac{1}{f'(\alpha_-)} \ln \ln M$ ; data by **Olivier Giraud**

## OTHER FACETS OF THE SAME STORY:

- **Statistics of high values of Riemann  $\zeta(1/2 + it)$**   
**YVF, G Hiary, J Keating**  
*Phys.Rev.Lett. 108 , 170601 (2012) & arXiv:1211.6063*
- **Fluctuations of the shape of Young diagrams**  
**sampled with the Plancherel measure. YF & S. Nechaev, in progress.**